

## 2 Brownian motion

### 2.1 Lévy's construction of BM

Consider the 1-d case, let  $B_t$  be the position (of a particle) at time  $t$ .

- (i) (start at origin)  $B_0 = 0$ , for convenience purposes;
- (ii) (independent increment)  $B_t - B_s$  is independent of the  $\sigma$ -algebra generated by  $\{B_r : r \leq s\}$ , for  $s < t$ ;
- (iii) (identically distributed increments)  $B_t - B_s$  has the same distribution as  $B_{t-s} - B_0$ , for  $s < t$ .
- (iv) (continuity w.p.1) the function  $t \mapsto B_t$  is continuous w.p.1

A process satisfying (i)-(iii) is a **Lévy process**. For instance, a Poisson process is a Lévy process but does not satisfy (iv)

**Proposition 2.1.** *If a process satisfies (i)-(iv), then there exists some  $\mu \in \mathbb{R}$  and  $\sigma^2 \geq 0$  s.t.  $B_t - B_s \sim N(\mu(t-s), \sigma^2(t-s))$ .*

A **Brownian motion** (or **Wiener Process**) with drift  $\mu$  and variance parameter  $\sigma^2$  starting at the origin is a stochastic process (a collection of random variables indexed by time  $t$ )  $\{B_t : t \geq 0\}$  satisfying (i), (ii), (iv), and  $B_t - B_s \sim N(\mu(t-s), \sigma^2(t-s))$  for  $s < t$ .

**Proposition 2.2.** *If  $B_t$  is a standard Brownian motion, then  $Y_t = \sigma B_t + t\mu$  is a Brownian motion with drift  $\mu$  and variance parameter  $\sigma^2$ .*

Does Brownian motion exist? We give a concrete construction.

**Lemma 2.3.** *Let  $X$  and  $Y$  be independent normal variables with zero mean and  $\sigma^2$  variance, then  $X + Y$  and  $X - Y$  are independent normal variables with zero mean and  $2\sigma^2$  variance.*

**Construction.** By Proposition 1.2, it suffices to construct a standard BM. It also suffices to consider only the interval  $[0, 1]$  by definition.

Consider the dyadic integers  $D_n = \{j2^{-n} : j = 0, 1, \dots, 2^n\}$  and  $D = \cup_{n \geq 0} D_n$ . We first construct  $B_t$  for  $t \in D$  and prove that it is uniformly continuous in  $t$  w.p.1. Since uniformly continuous functions defined on a dense set has a unique extension, we simply define  $B_t = \lim_{n \rightarrow \infty} B_{t_n}$  for  $t_n \rightarrow t$  and  $t_n \in D$ .

Let  $\{N_q : q \in D\}$  be a collection of independent standard normal random variables, then we define

$$\begin{cases} B_1 = N_1 \\ B_{1/2} = B_1/2 + N_{1/2}/2 \\ B_{1/4} = B_{1/2}/2 + N_{1/4}/2^{3/2} \\ B_{3/4} = 1 - B_{1/4} \\ \dots \\ B_d = B_{d+1/2^n}/2 + N_d/2^{(n+1)/2} \text{ for } d \in D_n \setminus D_{n-1} \end{cases} .$$

By Lemma 1.3,  $\{B_{1/2^n}, B_{2/2^n} - B_{1/2^n}, \dots, B_1 - B_{2^{n-1}/2^n}\}$  are independent  $N(0, 1/2^n)$  variables. We are left to show that  $q \mapsto B_1$  is uniformly continuous w.p.1.

It suffices to show that w.p.1, for any  $\alpha < 1/2$  and

$$K_n = \sup\{|B_q - B_s| : q, s \in D \text{ and } |q - s| \leq 1/2^n\},$$

$\lim_{n \rightarrow \infty} 2^{\alpha n} K_n = 0$  (in particular  $K_n \rightarrow 0$ ). Let

$$Y(j, n) = \sup\{|B_1 - B_{(j-1)/2^n}| : q \in D \cap [(j-1)/2^n, j/2^n]\}$$

and  $J_n = \max\{Y(j, n) : j = 1, \dots, 2^n\}$ , then  $J_n \leq 3K_n$ , and it suffices to substitute  $K_n$  for  $J_n$ . Observe that  $Y(1, n), \dots, Y(2^n, n)$  are i.i.d., and  $Y(1, n)$  has the same distribution as  $Y(1, 0)/2^{n/2}$  (the reason for  $2^{n/2}$  instead of  $2^n$  is because while the variance is halved, the distribution is scaled by a factor of  $\sqrt{2}$ ). Hence,

$$\begin{aligned} \mathbb{P}\{J_n \geq \epsilon_n\} &= \mathbb{P}\{\max Y(j, n) \geq \epsilon_n\} \\ &\leq \sum_{j=1}^{2^n} \mathbb{P}\{Y(j, n) \geq \epsilon_n\} \\ &= 2^n \mathbb{P}\{Y(1, n) \geq \epsilon_n\} \\ &= 2^n \mathbb{P}\{Y(1, 0) \geq 2^{n/2} \epsilon_n\}. \end{aligned}$$

By Borel-Cantelli lemma, if we can find  $\epsilon_n$  such that  $\sum_{n \geq 1} \mathbb{P}\{J_n \geq \epsilon_n\} < \infty$ , then  $J_n \leq \epsilon_n$  for large enough  $n$ . We need a lemma:

**Lemma 2.4** (reflection principle (for dyadics)). *For any  $a > 0$ ,  $\mathbb{P}\{Y(1, 0) > a\} \leq 4\mathbb{P}\{B_1 \geq a\}$ .*

*Proof.* Since  $\mathbb{P}\{Y(1, 0) > a\} \leq 2\mathbb{P}\{\sup_{q \in D} B_q \geq a\}$ , it suffices to show that  $\mathbb{P}\{\max_{j=1, \dots, 2^n} B_{j/2^n} > a\} \leq 2\mathbb{P}\{B_1 \geq a\}$  for any  $n$ . Fix  $n$  and let

$$A_j = \{B_{j/2^n} > a \text{ and } B_{i/2^n} \leq a \text{ for } i = 1, \dots, j-1\}$$

be the first time that  $B_{j/2^n}$  is larger than  $a$ , then  $\{\max_{j=1, \dots, 2^n} B_{j/2^n} > a\} = \sqcup_{k=1}^{2^n} A_k$ . Also observe that  $\mathbb{P}(B_1 > a | A_k) \geq 1/2$ . Hence

$$\mathbb{P}\{\max_{j=1, \dots, 2^n} B_{j/2^n} > a\} / 2 = \sum_{k=1}^{2^n} \mathbb{P}(A_k) / 2 \leq \sum_{k=1}^{2^n} \mathbb{P}\{B_1 > a; A_k\} = \mathbb{P}\{B_1 > a\}.$$

□

For large enough  $a$ , we have the following estimate:

$$\mathbb{P}\{Y(1, 0) > a\} \leq 4\mathbb{P}\{B_1 \geq a\} = 4 \int_a^\infty e^{-x^2/2} dx / \sqrt{2\pi} \leq 4 \int_a^\infty e^{-xa/2} dx / \sqrt{2\pi} \leq C e^{-a^2/2} / a \leq e^{-a^2/2}$$

where  $C$  is some constant. Hence we have

$$\mathbb{P}\{J_n \geq \epsilon_n 2^{-n/2}\} = 2^n \mathbb{P}\{Y(1, 0) \geq \epsilon_n\} \leq 2^n e^{-\epsilon_n^2/2} = e^{n(\log 2 - \beta^2/2)}$$

if we choose  $\epsilon_n = \beta\sqrt{n}$ . Hence if we choose  $\beta^2 > 2 \log 2$ , then  $\sum_{n \geq 1} \mathbb{P}\{J_n \geq \beta\sqrt{n} 2^{-n/2}\} < \infty$  as desired, and by Borel-Cantelli lemma  $J_n \leq \beta\sqrt{n} 2^{-n/2}$ . Hence  $\lim_{n \rightarrow \infty} 2^{\alpha n} J_n = 0$  for  $\alpha < 1/2$ .

Finally, one can check that  $B_t = \lim_{n \rightarrow \infty} B_{t_n}$  for  $t_n \rightarrow t$  and  $t_n \in D$  does define a Brownian motion.

**Remark.** We need our probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  to be large enough to have a countable collection of independent standard normal random variables. Luckily,  $[0, 1]$  with Borel  $\sigma$ -field and Lebesgue measure suffices (HW1).

## 2.2 Properties of BM

**Proposition 2.5.** *W.p.1, Brownian motion  $\{B_t\}_{0 \leq t \leq 1}$  is not differentiable at  $t$  for all  $t \in [0, 1]$ .*

*Proof.* Suppose event  $E_M = \{\exists t \in [0, 1] \text{ such that } |B'_t| \leq M\}$  holds, then by continuity there exists  $\epsilon$  s.t.  $|s - t| < \epsilon$  implies  $|B_t - B_s| < 2M\epsilon$ . Then if  $|s - t|, |s' - t| < \epsilon$ , then  $|B_s - B_{s'}| \leq 4M\epsilon$ . For large enough  $n$ , this means that there exists  $k \leq n$  s.t.

$$A_{M,n} \begin{cases} |B_{(k+1)/n} - B_{k/n}| \leq 4M/n \\ |B_{(k+2)/n} - B_{(k+1)/n}| \leq 4M/n \\ |B_{(k+3)/n} - B_{(k+2)/n}| \leq 4M/n \end{cases} .$$

To show that the countable union  $\mathbb{P}(\cup_{M \geq 1} E_M) = 0$ , it suffices to show that  $\mathbb{P}(E_M) = 0$ . Observe that  $\mathbb{P}(E_M) \leq \mathbb{P}\{\exists k \leq n : A_{M,n}\}$ , so it suffices to show that  $\lim_{n \rightarrow \infty} \mathbb{P}(A_{M,n}) = 0$ . Exactly,

$$\begin{aligned} \mathbb{P}(A_{M,n}) &= \mathbb{P} \left\{ \min_{0 \leq k \leq n-1} \left[ \max \begin{cases} |B_{(k+1)/n} - B_{k/n}|, \\ |B_{(k+2)/n} - B_{(k+1)/n}|, \\ |B_{(k+3)/n} - B_{(k+2)/n}| \end{cases} \leq 4M/n \right] \right\} \\ &\leq \sum_{k=0}^{n-1} \mathbb{P} \left\{ \max \begin{cases} |B_{(k+1)/n} - B_{k/n}|, \\ |B_{(k+2)/n} - B_{(k+1)/n}|, \\ |B_{(k+3)/n} - B_{(k+2)/n}| \end{cases} \leq 4M/n \right\} \\ &= n \mathbb{P} \left\{ \max\{|B_{1/n}|, |B_{2/n} - B_{1/n}|, |B_{3/n} - B_{2/n}|\} \leq 4M/n \right\} \\ &= n [\mathbb{P}\{|B_{1/n}| \leq 4M/n\}]^3 \\ &= n [\mathbb{P}\{|\sqrt{n}B_{1/n}| \leq 4M/\sqrt{n}\}]^3 \\ &= n \left[ \int_0^{4M/\sqrt{n}} e^{-x^2/2} dx / \sqrt{2\pi} \right]^3 \leq CM^3 n / (\sqrt{n})^3 \rightarrow 0 \end{aligned}$$

as  $n \rightarrow \infty$ . □

**Remark.** Proposition 1.5 is stronger than: for any  $t \in [0, 1]$ ,  $B_t$  is not differentiable at  $t$  w.p.1.

The **natural filtration** for Brownian motion is  $\mathcal{F}_t = \sigma\{B_s : s \leq t\}$ . We say a process  $B_t$  is adapted to a filtration  $\{\mathcal{F}_t : t \geq 0\}$  if for  $B_t$  is  $\mathcal{F}_t$ -measurable for each  $t$ . Recall that a **martingale** w.r.t.  $\mathcal{F}_t$  is an  $\mathcal{F}_t$ -adapted process such that  $\mathbf{E}(|M_t|) < \infty$  for any  $t$  and  $\mathbf{E}(M_t | \mathcal{F}_s) = M_s$  for any  $s < t$ . A martingale is continuous (not merely continuous-time!) if  $t \mapsto M_t$  is a continuous function w.p.1. Observe that the *standard* Brownian motion is a continuous martingale:

$$\mathbf{E}(B_t | \mathcal{F}_s) = \mathbf{E}(B_s + (B_t - B_s) | \mathcal{F}_s) = \mathbf{E}(B_s | \mathcal{F}_s) + \mathbf{E}((B_t - B_s) | \mathcal{F}_s) = B_s + \mathbf{E}(B_t - B_s) = B_s.$$

A random variables  $T : \Omega \rightarrow [0, \infty]$  is a stopping time w.r.t.  $\{\mathcal{F}_t\}$  if  $\{T \leq t\} \in \mathcal{F}_t$  for all  $t \geq 0$ . Also define  $\mathcal{F}_T = \sigma\{A : A \cap \{T \leq t\} \in \mathcal{F}_t \text{ for all } t\}$ .

**Proposition 2.6** (reflection principle). *Let  $B_t$  be a Brownian motion with drift  $\mu = 0$ , then for any  $a > 0$ ,  $\mathbb{P}\{\max_{0 \leq s \leq t} B_s \geq a\} = 2\mathbb{P}\{B_t \geq a\}$ .*

*Proof.* Let  $T_a = \min\{s : B_s \geq a\} = \min\{s : B_s = a\}$ , then  $\mathbb{P}\{\max_{0 \leq s \leq t} B_s \geq a\} = \mathbb{P}\{T_a \leq t\} = \mathbb{P}\{T_a < t\}$ . Also  $\mathbb{P}\{B_t \geq a | t > T_a\} = 1/2$ , hence

$$\mathbb{P}\{B_t \geq a\} = \mathbb{P}\{T_a < t\} \mathbb{P}\{B_t \geq a | T_a < t\} = \mathbb{P}\{T_a < t\} / 2.$$

□

**Proposition 2.7** (weak Markov property). *Let  $B_t$  be a Brownian motion and  $Y_s = B_{s+t} - B_t$ , then  $\{Y_s : s \geq 0\}$  is a Brownian motion with the same  $\mu$  and  $\sigma^2$  as  $B_t$ , and is independent of  $\mathcal{F}_t = \sigma\{B_s : s \leq t\}$ .*

**Proposition 2.8** (strong Markov property). *Let  $B_t$  be a  $BM(\mu, \sigma^2)$  and  $\tau$  a stopping time w.r.t the filtration  $\{\sigma(B_t)\}$ . Suppose that  $\mathbb{P}\{\tau < \infty\} = 1$ . Then  $Y_t = B_{t+\tau} - B_\tau$  is a  $BM(\mu, \sigma^2)$  independent of  $\sigma\{B_t : t \in [0, \tau]\}$ .*

**Proposition 2.9.** *Let  $B_t$  be a Brownian motion and  $T$  be a stopping time all w.r.t.  $\{\mathcal{F}_t\}$ . Suppose that  $\mathbb{P}\{T \leq \infty\} = 1$ . Let  $Y_t = B_{t+T} - B_T$  for  $0 \leq t < \infty$ . Then  $Y_t$  is a Brownian motion with the same  $\mu$  and  $\sigma^2$ , and is independent of  $\mathcal{F}_t$ .*

**Example.** We wish to find  $q(r, s) = \mathbb{P}\{B_t = 0 \text{ for some } t \in [r, s]\}$  where  $B_t$  is a standard Brownian motion.

**Lemma 2.10.** *Let  $A_1, A_2, \dots$  be events and  $\mathcal{F}_n = \sigma\{A_1, \dots, A_n\}$ . Suppose that there exists  $b_n > 0$  such that  $\sum_{n \geq 1} b_n = \infty$  and  $\mathbb{P}(A_{n+1} | \mathcal{F}_n) \geq b_n$  for any  $n$ . Then  $A_n$  occurs infinitely often.*

**Remark.** This lemma generalizes the Borel-Cantelli lemma: if  $A_1, A_2, \dots$  are independent events and  $\sum_{n \geq 1} \mathbb{P}(A_n) = \infty$  then  $A_n$  occurs infinitely often.

**Proposition 2.11** (law of iterated logarithm). *Let  $B_t$  be a standard BM. Then w.p.1,*

$$\limsup_{t \rightarrow \infty} B_t / \sqrt{2t \log \log t} = 1.$$

*Proof.* We first show that  $\limsup_{t \rightarrow \infty} B_t / \sqrt{t \log \log t} \leq \sqrt{2}$  w.p.1. It suffices to show that for any  $\epsilon > 0$  w.p.1 for all  $t$  sufficiently large,  $B_t \leq \sqrt{2(1+\epsilon)t \log \log t}$ . We use a trick called geometric scaling. Let  $\rho > 1$  and  $A_n^{\epsilon, \rho} = \{B_t \geq \sqrt{2(1+\epsilon)t \log \log t} \text{ for some } \rho^{n-1} \leq t \leq \rho^n\}$ . By Borel-Cantelli lemma, it suffices to show that for any  $\epsilon > 0$  there exists  $\rho > 1$  such that  $\sum_{n \geq 1} \mathbb{P}(A_n^{\epsilon, \rho}) < \infty$ . By the reflection principle,

$$\begin{aligned} \mathbb{P}(A_n^{\epsilon, \rho}) &\leq \mathbb{P}\left\{ \sup_{0 \leq t \leq \rho^n} B_t \geq \sqrt{2(1+\epsilon)\rho^{n-1} \log \log \rho^{n-1}} \right\} \\ &\leq 2\mathbb{P}\left\{ B_{\rho^n} \geq \sqrt{2(1+\epsilon)\rho^{n-1} \log \log \rho^{n-1}} \right\} \\ &= 2\mathbb{P}\left\{ B_{\rho^n} / \sqrt{\rho^n} \geq \sqrt{2(1+\epsilon)(\log(n-1) + \log \log \rho) / \rho} \right\} = C_\rho (n-1)^{-(1+\epsilon)/\rho} \end{aligned}$$

for some constant  $C_\rho$  depending on  $\rho$ . It suffices to take  $\rho$  so that  $(1+\epsilon)/\rho > 1$  to make it summable.

For the other side of the inequality, we will show that for any  $\epsilon > 0$  there exists  $\rho > 1$  such that  $\limsup_{n \rightarrow \infty} B_{\rho^n} / \sqrt{2(1-\epsilon)\rho^n \log \log \rho^n} \geq 1$  w.p.1. Let  $A_n = \{|B_{\rho^n}| \geq \sqrt{2(1-\epsilon)\rho^n \log \log \rho^n}\}$ . Since  $\mathbb{P}(A_n | \mathcal{F}_{n-1}) \geq \mathbb{P}\{B_{\rho^n} - B_{\rho^{n-1}} \geq \sqrt{2(1-\epsilon)\rho^n \log \log \rho^n}\} / 2 = \mathbb{P}(\tilde{A}_n) / 2$ , where  $\mathcal{F}_{n-1} = \sigma\{A_1, \dots, A_{n-1}\}$ , by the previous lemma it suffices to show that  $\sum_{n \geq 1} \mathbb{P}(\tilde{A}_n) = \infty$ . Observe that

$$\begin{aligned} \mathbb{P}(\tilde{A}_n) &= \mathbb{P}\left\{ (B_{\rho^n} - B_{\rho^{n-1}}) / \sqrt{\rho^n - \rho^{n-1}} \geq \sqrt{2(1-\epsilon)\rho(\log n + \log \log \rho) / (\rho - 1)} \right\} \\ &\geq C_{\epsilon, \rho} \exp\left\{ -(1-\epsilon)\rho(\log n + \log \log \rho) / (\rho - 1) \right\} / \sqrt{2(1-\epsilon)\rho(\log n + \log \log \rho) / (\rho - 1)} \\ &\geq C'_{\epsilon, \rho} n^{-(1-\epsilon)\rho / (\rho - 1)} / \sqrt{\log n}. \end{aligned}$$

Again, take  $\rho_\epsilon$  large enough so that  $(1-\epsilon)\rho / (\rho - 1) < 1$  to make unsummable.  $\square$

**Remark.** The common strategy in proving these statements is to use Borel-Cantelli lemma first. We often need some scaling and to investigate the difference of Brownian paths  $(B_{\rho^n} - B_{\rho^{n-1}})$ . Then to bound the probability of individual events, use tools like reflection principle and take integration of the Gaussian variable.

Let  $B_t$  be a standard BM, the **zero set** of it is  $Z_t = \{s \in [0, t] : B_s = 0\}$ . We write  $Z = Z_\infty$ . A point  $s \in Z_t$  is **right-isolated** if there exists  $\epsilon > 0$  such that  $r \notin Z$  for  $r \in (s, s + \epsilon)$ . A point in  $Z_t$  is **isolated** if it is both left- and right-isolated.

**Proposition 2.12.** *We make the following observations:*

- (i)  $0 \in Z_t$ ;
- (ii)  $Z_t$  is closed for any  $t$ ;
- (iii)  $0$  is not right isolated;
- (iv)  $Z$  is unbounded;
- (v)  $Z$  has no isolated points.

*Proof.* (i)-(iii) is immediate. To prove (iv), we use intermediate value theorem together with, say, the law of iterated logarithm. Alternatively we can use the inverse ....

For (v), □

## 2.3 Quadratic variation of BM

## 2.4 Dimension of BM

## 2.5 BM in $\mathbb{R}^d$

**Definition.** Let  $B_t^1, \dots, B_t^d$  be a sequence of independent (1-dimensional) standard BMs, then  $B_t = (B_t^1, \dots, B_t^d)$  is a standard  $d$ -dimensional BM.

**Lemma 2.13.** *For standard  $d$ -dimensional BM  $B_t$ , the following holds:*

- (i)  $B_t = 0$ ;
- (ii) if  $s < t$ , then  $B_t - B_s$  is independent of  $\mathcal{F}_s$ ;
- (iii)  $(B_t - B_s) \sim N(0, (t - s)I)$  where  $I$  is the identity matrix, and the density is given by

$$f(x_1, \dots, x_n) = \prod_{j=1}^d (2\pi(t - s))^{-1/2} \exp\left\{\frac{-x_j^2}{2(t - s)}\right\} = (2\pi(t - s))^{-d/2} \exp\left\{\frac{-|x^2|}{2(t - s)}\right\};$$

if  $B_t$  has drift  $\mu$  and covariance matrix  $\Gamma$ , then  $(B_t - B_s) \sim N((t - s)\mu, (t - s)\Gamma)$  instead;

- (iv) w.p.1  $t \mapsto B_t$  is continuous;
- (v)  $Y_t = \mu t + AB_t$  is a BM with drift  $\mu$  and covariance matrix  $AA^T$  for  $\mu \in \mathbb{R}^d$  and  $A \in M_d(\mathbb{R})$ .

## 2.6 BM as martingales

Let  $M_t$  be a continuous martingale and  $\tau$  a stopping time, both adapted to the filtration  $\{\mathcal{F}_t\}$ . Let  $\hat{M}_t = M_{t \wedge \tau}$  and suppose that  $\mathbb{P}\{\tau < \infty\} = 1$ . Then  $\hat{M}_\infty = M_\tau$ , and  $\mathbf{E}(M_0) = \mathbf{E}(\hat{M})$  for every  $t$ . The *optional stopping theorem* says that if  $\lim_{t \rightarrow \infty} \mathbf{E}(\hat{M}_t) = \mathbf{E}(\hat{M}_\infty)$ , then  $\mathbf{E}(M_\tau) = \mathbf{E}(M_0)$ .

**Example** (gambler's ruin for BM). Let  $B_t$  be a standard BM and  $\tau = \inf\{t : B_t = a \text{ or } B_t = -b\}$  be the stopping time. Then  $\mathbb{P}\{T < \infty\} = 1$  by the recurrence property of BM, and  $\mathbf{E}(B_0) = 0$ . Since  $\mathbf{E}(B_\tau) = \lim_{t \rightarrow \infty} \mathbf{E}(B_{t \wedge \tau})$  by the dominating convergence theorem, we get  $0 = \mathbf{E}(B_\tau) = a\mathbb{P}\{B_\tau = a\} - b\mathbb{P}\{B_\tau = -b\}$  which solves to  $\mathbb{P}\{B_\tau = a\} = b/(a+b)$  and  $\mathbb{P}\{B_\tau = -b\} = a/(a+b)$ .

## 2.7 Harmonic functions in $\mathbb{R}^d$

Let  $D$  be a connected open subset of  $\mathbb{R}^d$  and  $z \in D$  a point with  $\text{dist}(z, \partial D) > \epsilon$ . Define the **(spherical) mean value** of  $f$  on the sphere  $B_\epsilon(z)$  of radius  $\epsilon$  about  $z$ :

$$\text{MV}(f, z, \epsilon) = \int_{|\omega - z| = \epsilon} f(\omega) ds(\omega)$$

where  $s$  is the surface measure (constant times the surface area) normalized s.t.  $\text{MV}(1, z, \epsilon) = 1$ . A function  $f : D \rightarrow \mathbb{R}$  is **harmonic** if it is continuous and satisfies the *mean value property*:  $\text{MV}(f, z, \epsilon) = f(z)$  for every  $z \in D$  with  $\text{dist}(z, \partial D) > \epsilon$ .

Let  $\Delta f = \sum_{j=1}^d \partial^2 f / \partial x_j^2$  be the *Laplacian* of  $f$ , then a function  $f : D \rightarrow \mathbb{R}$  is harmonic iff it is  $C^2$  and  $\Delta f(z) = 0$  for all  $z \in D$ . When dimension  $d = 1$ , then harmonic is equivalent to linear, but when  $d > 1$  the class of harmonic functions is larger.

## 2.8 Dirichlet problem

The Dirichlet problem asks: let  $D \subset \mathbb{R}^d$  be a bounded domain and  $F : \partial D \rightarrow \mathbb{R}$  a continuous function, can we find an extension  $f : \bar{D} \rightarrow \mathbb{R}$  of  $F$  such that:

- (i)  $f = F$  on  $\partial D$ ;
- (ii)  $f$  is continuous on  $\bar{D}$ ;
- (iii)  $\Delta f(x) = 0$  for all  $x \in D$ .

We first show that if exists, such  $f$  is unique:

Let  $B_t$  be a  $d$ -dimensional BM and  $\tau = \inf\{t \geq 0 : B_t \in \partial D\}$ . Define  $f(x) = \mathbf{E}^x[F(B_\tau)]$ , then  $f(x)$  is harmonic and continuous on  $D$ , and satisfies  $f = F$  on  $\partial D$ . The only thing missing from  $f$  being a solution to the Dirichlet problem is the continuity on  $\partial D$ . It turns out,  $f$  is only continuous on all the regular points  $z \in \partial D$ .

Let  $X_t$  be a (time-homogeneous) Markov process taking values in  $\mathbb{R}^d$ , the **infinitesimal generator**  $L$  is the operator is defined by

$$Lg(x) = \lim_{t \downarrow 0} \frac{\mathbf{E}^x[g(X_t)] - g(x)}{t}.$$

**Proposition 2.14.** *Let  $B_t$  be a standard BM in  $\mathbb{R}^d$  and  $f$  a  $C^2$  function (that does not grow too fast at  $\infty$ ), then  $Lf(x) = \Delta f(x)/2$ .*

*Proof.* WLOG assume that  $x = 0$  and  $f(0) = 0$ . Since  $f$  is  $C^2$ , the Taylor expansion gives

$$f(B_t) = \sum_{j=1}^d \partial_j f(0) B_t^j + \frac{1}{2} \sum_{j=1}^d \sum_{k=1}^d \partial_{jk} f(0) B_t^j B_t^k + o(|B_t|^2).$$

Hence  $Lf(x) = \lim_{t \downarrow 0} (\mathbf{E}^x[f(B_t)] - f(x))/t = ?$ . □

**Remark.** If the BM  $B_t$  has mean  $\mu \in \mathbb{R}^d$  and covariance matrix  $\Gamma$ , then

$$Lf(x) = \sum_{j=1}^d \mu_j \partial_j f(x) + \frac{1}{2} \sum_{j=1}^d \sum_{k=1}^d \partial_{jk} f(x) \Gamma_{jk}$$

**Proposition 2.15.** *Let  $B_t$  be a standard  $d$ -dimensional BM and  $f : \mathbb{R}^d \rightarrow \mathbb{R}$  be bounded. Then for  $t > 0$ ,  $\varphi(t, x) = \mathbf{E}^x[f(B_t)]$  satisfies the heat equation*

$$\partial_t \varphi(t, x) = \frac{1}{2} \Delta_x \varphi(t, x).$$

*Proof.* A Brownian motion  $B_t$  starting at  $x$  has density  $p_t(x, y) = (2\pi t)^{-d/2} e^{-|y-x|^2/2t}$ , where  $\partial_t p_t = \Delta_x p_t(x, y)/2$ . Since  $f$  is bounded, we can exchange integral and derivative:

$$\begin{aligned} \partial_t \varphi(t, x) &= \partial_t \int_{\mathbb{R}^d} f(y) p_t(x, y) dy \\ &= \int_{\mathbb{R}^d} f(y) \partial_t p_t(x, y) dy \\ &= \frac{1}{2} \int_{\mathbb{R}^d} f(y) \Delta_x p_t(x, y) dy \\ &= \frac{1}{2} \Delta_x \int_{\mathbb{R}^d} f(y) p_t(x, y) dy = \frac{1}{2} \Delta_x \varphi(t, x). \end{aligned}$$

□

Let  $B_t$  be a standard  $d$ -dimensional BM. For  $d \geq 3$ , the **Green's function**  $G(x, y)$  is defined by

$$\begin{aligned} G(x, y) &= \lim_{\epsilon \downarrow 0} \mathbf{E}^x \left[ \frac{1}{V_\epsilon} \int_0^\infty \mathbf{1}_{\{|B_t - y| \leq \epsilon\}} dt \right] \\ &= \lim_{\epsilon \downarrow 0} \frac{1}{V_\epsilon} \int_0^\infty \int_{|y-z| \leq \epsilon} p_t(x, z) dz dt \\ &= \int_0^\infty p_t(x, y) dt \\ &= \int_0^\infty \frac{1}{(2\pi t)^{d/2}} \exp\left\{ -\frac{|y-x|^2}{2t} \right\} dt = \frac{\Gamma(d/2 - 1)}{2\pi^{d/2}} |x-y|^{2-d}, \end{aligned}$$

where  $V_\epsilon$  is the ( $d$ -dimensional) volume of the  $\epsilon$ -ball. Intuitively,  $G(x, y)$  represents the expected amount of times spent in  $y$  of a standard Brownian motion starting at  $x$ .  $g_y(x) = G(x, y)$  is harmonic for  $x \neq y$ . Also define

$$G(x) = \int_0^\infty \frac{1}{(2\pi t)^{d/2}} \exp\left\{ -\frac{|x|^2}{2t} \right\} dt$$

and  $G_D(x, y)$  to be the expected amount of time spent in  $y$  before leaving  $D$  of a standard BM starting at  $x$ . Then  $G(x, y) = G_D(x, y) + \mathbf{E}^x[G(B_\tau, y)]$  where  $\tau = \inf\{t : B_t \in \partial D\}$ .

**Proposition 2.16.** *Let  $g(y) = G_D(x, y)$ , then*

- (i)  $g$  is harmonic on  $D \setminus \{x\}$ ;
- (ii)  $g(y) \rightarrow 0$  for  $y \rightarrow z \in \partial D$  and  $z$  regular;
- (iii)  $G_D(x, y) = G_D(y, x)$ .

For  $d \leq 2$ , the recurrence property of BM makes  $G(x) = \infty$ . Hence we redefine, for  $d = 2$ , the **potential kernel**

$$G(x) = \int_0^\infty [p_t(0, z) - p_t(o, x)] dt = \int_0^\infty \frac{1}{2\pi t} \left[ e^{-1/2t} - e^{-|x|^2/2t} \right] dt = -\frac{1}{\pi} \log |x|$$

where we take  $z = (0, 1)$ . This is a radially symmetric harmonic function on  $\mathbb{R}^2 \setminus \{0\}$ , and carries the intuitive meaning of “time spent in  $x$  compared to the time spent on the unit circle. Hence  $G(x) = 0$  for  $|x| = 1$ . When  $t$  is large,  $e^{-1/2t} - e^{-|x|^2/2t} = \frac{|x|^2 - 1}{2t} + O(1/t^2)$ , hence  $G(x) < \infty$  for  $x \neq 0$ . If we choose  $z \neq (0, 1)$ , then  $G(x)$  will differ from  $-\log |x|/\pi$  by a constant.

Proposition 2.16 still works for this  $G_D(x, y)$  defined as such in  $d = 2$ , where

$$G_D(x, y) = \mathbf{E}^x \left[ \frac{\log |B_\tau - z|}{\pi} \right] - \frac{\log |x - z|}{\pi}.$$

Let  $D \subset \mathbb{R}^2$  be a bounded domain. Let  $G_D(x, y) = \int_0^\infty p_t^D(x, y) dt$  where  $p_t^D(x, y)$  is the density of BM starting at  $x$  and killed when reaching the boundary, then for  $f : D \rightarrow \mathbb{R}$  a continuous function,

$$\mathbf{E}^x [f(B_t) \mathbf{1}_{\{t < T\}}] = \int_D p_t^D(x, y) f(y) dy.$$

Also,  $p_t^D(x, y)$  satisfies the heat equation:  $\partial_t p_t^D(x, y) = \Delta_x p_t^D(x, y)/2$ , and  $p_t^D(x, y) = p_t^D(y, x)$ . We say that  $p_t^D(x, y)$  is the solution to the heat equation with initial condition  $p_0(x, y) = \delta_{xy}$ . Also

$$p_t^D(x, y) = p_t(x, y) - \mathbf{E}^x [p_{t-\tau}(B_\tau, y) \mathbf{1}_{\{t > T\}}].$$



### 3 Stochastic Calculus

#### 3.1 Construction

Consider the differential equation

$$dX_t = R_t dt + A_t dB_t$$

where  $B_t$  is a BM, and  $X_t$  and  $R_t$  are random variables. This equation carries the meaning “at time  $t$ ,  $X_t$  looks *locally* like a  $\text{BM}(R_t, A_t^2)$ ”. We can not differentiate both sides by  $dt$  since BM is non-differentiable. But we can consider

$$X_t = X_0 + \int_0^t R_s ds + \int_0^t A_s dB_s.$$

It remains to make sense of the last term.

A  $\{\mathcal{F}_t\}$ -adapted-process  $A_t$  is **simple** if there exists  $0 = t_0 < t_1 < \dots < t_n = \infty$  and  $L^2$ -random variables  $Y_0, \dots, Y_n$  such that  $Y_j$  is  $\mathcal{F}_{t_j}$ -measurable and  $A_t = Y_j$  for  $t \in [t_j, t_{j+1})$ .

Let  $B_t$  be a standard BM and  $A_t$  be a  $\mathcal{F}_t$ -adapted standard BM. The (Itô) **stochastic integral** of  $A_t$  is defined by

$$Z_t = \int_0^t A_s dB_s = \sum_{i=0}^{j-1} Y_i (B_{t_{i+1}} - B_{t_i}) + Y_j (B_t - B_{t_j})$$

for  $t \in [t_j, t_{j+1}]$ .

**Proposition 3.1.** *The stochastic integral defined satisfies the following properties:*

(i)  $Z_t$  is adapted to  $\{\mathcal{F}_t\}$ ;

(ii) (linearity) let  $a$  and  $b$  be constants, and  $C_t$  be a simple process, then

$$\int_0^t (aA_s + bC_s) dB_s = a \int_0^t A_s dB_s + b \int_0^t C_s dB_s;$$

also if  $r \in [0, t]$ , then

$$\int_0^t A_s dB_s = \int_0^r A_s dB_s + \int_r^t A_s dB_s;$$

(iii) (martingale)  $Z_t$  is an  $L^2$ -martingale;

(iv) (variance isometry)

$$\text{Var}(Z_t) = \mathbf{E}(Z_t^2) = \int_0^t \mathbf{E}(A_s^2) dB_s;$$

(v) (continuity) w.p.1,  $t \mapsto Z_t$  is a continuous function.

*Proof.* □

**Proposition 3.2.** *Let  $A_t$  be a bounded, continuous, and  $\{\mathcal{F}_t\}$ -adapted process. Then for any  $t_0$ , there exists a sequence of bounded simple processes  $A_t^{(n)}$  converging to  $A_t$  in the sense that*

$$\lim_{n \rightarrow \infty} \mathbf{E} \left[ \int_0^{t_0} (A_t - A_t^{(n)})^2 dt \right] = 0.$$

Hence we extend the definition of stochastic integral for bounded continuous (adapted) process  $A_t$ :

$$Z_t = \int_0^t A_s dB_s = \lim_{n \rightarrow \infty} \int_0^t A_s^{(n)} dB_s.$$

**Remark.** properties, continuity

A  $\{\mathcal{F}_t\}$ -adapted process  $A_s(\omega)$  is **progressively measurable** if  $A_s(\omega)$  is measurable in  $\mathcal{B}([0, t]) \otimes \mathcal{F}_t$  measurable for  $s \in [0, t]$ .

### 3.2 Itô's formula

**Proposition 3.3** (Itô's formula). *Let  $B_t$  be a standard BM and  $f : \mathbb{R} \rightarrow \mathbb{R}$  be  $C^2$ , then for every  $t$ ,*

$$f(B_t) - f(B_0) = \int_0^t f'(B_s) dB_s + \frac{1}{2} \int_0^t f''(B_s) ds. \quad (*)$$

*Proof.* Since both sides of (\*) are continuous, it suffices to show (\*) for dyadics  $t \in \mathfrak{D}$ . We do the case for  $t = 1$ , the other  $t \in \mathfrak{D}$  will be similar. Let  $0 = t_0 < t_1 < \dots < t_n = 1$  be a partition. Then

$$f(B_t) - f(B_0) = \sum_{j=1}^n (f(B_{t_j}) - f(B_{t_{j-1}})).$$

Assume that there exists  $C < \infty$  such that  $|f''| < C$ . By Taylor expansion, we have

$$\frac{1}{2} m_{j,n} (B_{t_j} - B_{t_{j-1}})^2 \leq |f(B_{t_j}) - f(B_{t_{j-1}}) - f'(B_{t_{j-1}})(B_{t_j} - B_{t_{j-1}})| \leq \frac{1}{2} M_{j,n} (B_{t_j} - B_{t_{j-1}})^2$$

where  $M_{j,n} = \sup_{z \in [B_{t_{j-1}}, B_{t_j}]} f''(z)$  and  $m_{j,n} = \inf_{z \in [B_{t_{j-1}}, B_{t_j}]} f''(z)$ . Refine the partition so that

$$\begin{aligned} \sum_{j=1}^n f'(B_{t_{j-1}})(B_{t_j} - B_{t_{j-1}}) &\rightarrow \int_0^1 f'(B_s) dB_s, \text{ and} \\ \sum_{j=1}^n \frac{1}{2} m_{j,n} (B_{t_j} - B_{t_{j-1}})^2, \sum_{j=1}^n \frac{1}{2} M_{j,n} (B_{t_j} - B_{t_{j-1}})^2 &\rightarrow \frac{1}{2} \int_0^1 f''(B_s) ds. \end{aligned}$$

If  $|f|$  is not bounded, then let  $T_K = \inf\{t : |f''(B_t)| \geq K\}$ . Then Itô's formula holds for  $t \wedge T_K$ . Now let  $K \rightarrow \infty$ .  $\square$

**Remark.** Itô's formula is often written in the differential form:

$$df(B_t) = f'(B_t) dB_t + \frac{1}{2} f''(B_t) dt.$$

Intuitively, the process  $Y_t = f(B_t)$  evolves like a Brownian motion with drift  $f'(B_t)/2$  and variance  $f''(B_t)^2$  locally at time  $t$ .

**Proposition 3.4** (Itô's formula, II). *Let  $f(t, x)$  be  $C^1$  in  $t$  and  $C^2$  in  $x$ , then*

$$f(t, B_t) - f(0, B_0) = \int_0^t \left( \partial_t f(s, B_s) + \frac{1}{2} \partial_x^2 f(s, B_s) \right) ds + \int_0^t \partial_x f(s, B_s) dB_s$$

*Proof.* Again, it suffices to show for  $t = 1$ . Observe that

$$f(t_j, B_{t_j}) - f(t_{j-1}, B_{t_{j-1}}) = f(t_j, B_{t_j}) - f(t_{j-1}, B_{t_j}) + f(t_{j-1}, B_{t_j}) - f(t_{j-1}, B_{t_{j-1}}).$$

Hence taking sum and refining the partition

$$\sum_{j=1}^n f(t_j, B_{t_j}) - f(t_{j-1}, B_{t_{j-1}}) = \sum_{j=1}^n \partial_t f(t_j, B_{t_j})(t_j - t_{j-1}) + \dots$$

□

Let  $f(t, x) = e^{at+bx}$  and  $X_t = f(t, B_t)$ , then

$$dX_t = X_t \left[ \left( a + \frac{b^2}{2} \right) dt + b dB_t \right].$$

Hence solution to the SDE  $dX_t = X_t(\mu dt + b dB_t)$  is the **geometric Brownian motion**

$$X_t = X_0 \exp\{(\mu - \sigma^2/2)t + \sigma B_t\}.$$

**Proposition 3.5.** *Let  $\{X_t\}_{0 \leq t \leq 1}$  be a process satisfying  $dX_t = R_t dt + A_t dB_t$ , then the quadratic variation*

$$\langle X \rangle_t = \lim_{n \rightarrow \infty} \sum_{j=1}^n (X_{jt/n} - X_{(j-1)t/n})^2$$

of  $X_t$  is given by

$$\left\langle \int_0^t A_s dB_s \right\rangle = \lim_{n \rightarrow \infty} \sum_{j=1}^n \left( \int_0^{tj/n} A_s dB_s \right)^2 = \lim_{n \rightarrow \infty} \sum_{j=1}^n \int_0^{tj/n} A_s^2 dB_s.$$

We write  $d\langle X \rangle_t = A_t^2 dt$ .

*Proof.* Let  $j = 1$ , then

$$\begin{aligned} (X_{t/n} - X_0)^2 &= \left( \int_0^{t/n} R_s ds + \int_0^{t/n} A_s dB_s \right)^2 \\ &= \left( \int_0^{t/n} R_s ds \right)^2 + 2 \left( \int_0^{t/n} R_s ds \right) \left( \int_0^{t/n} A_s dB_s \right) + \left( \int_0^{t/n} A_s dB_s \right)^2 \\ &= O(1/n^2) + O(1/n)O(1/\sqrt{n}) + O(1/n). \end{aligned}$$

There are  $n$  terms of this form in the sum. Hence as  $n \rightarrow \infty$  the only term contributing to the sum is the last term of order  $O(1/n)$ . □

**Proposition 3.6** (Itô's formula (III)). *Let  $X_t$  be the process that satisfies  $dX_t = R_t dt + A_t dB_t$ . Let  $f(t, x)$  be  $C^1$  in  $t$  and  $C^2$  in  $x$ . Then*

$$\begin{aligned} df(t, X_t) &= \partial_t f(t, X_t) dt + \partial_x f(t, X_t) dX_t + \frac{1}{2} \partial_{xx}^2 f(t, X_t) d\langle X \rangle_t \\ &= \left( \partial_t f(t, X_t) + R_t \partial_x f(t, X_t) + \frac{1}{2} A_t^2 \partial_{xx}^2 f(t, X_t) \right) dt + A_t \partial_x f(t, X_t) dB_t. \end{aligned}$$

## Diffusion

A **diffusion process**  $X_t$  is a process satisfying

$$dX_t = m(t, X_t) dt + \sigma(t, X_t) dB_t \quad (*)$$

with certain initial condition  $X_0 = x_0$ . We can simulate a diffusion process using stochastic Euler's method: take  $x_0 = 0$  and small  $\Delta t$  and define recursively

$$X_{(k+1)\Delta t} = X_{k\Delta t} + m(k\Delta t, X_{k\Delta t})\Delta t + \sigma(k\Delta t, X_{k\Delta t})\sqrt{\Delta t}N,$$

where  $N \sim N(0, 1)$ . Suppose  $f$  is nice enough (exchanging limit and integral) so that the generator

$$Lf(x) = \lim_{t \downarrow 0} \frac{\mathbf{E}^x[f(X_t)] - f(x)}{t}$$

is defined. For simplicity assume time homogeneity ( $m(t, x) = m(x)$  and  $\sigma(t, x) = \sigma(x)$ ) with  $\sigma$  and  $m$  both continuous and bounded. By Itô's formula

$$\mathbf{E}^x[f(X_t)] = \mathbf{E} \left[ \int_0^t \left( m(X_s)f'(X_s) + \frac{1}{2}\sigma^2(X_s)f''(X_s) \right) ds + \int_0^t \sigma(X_s)f'(X_s) dB_t \right].$$

The last term is a martingale with expectation 0. Hence the generator is given by

$$Lf(x) = m(x)f'(x) + \frac{1}{2}\sigma^2(x)f''(x).$$

Now we want to solve the SDE given by (\*) with initial condition  $X_0 = x_0$  to show that diffusion process  $X_t$  actually exists, under the assumption that  $m$  and  $\sigma$  are uniformly Lipschitz in  $x$ :

$$|m(t, x) - m(t, y)| \leq \beta|x - y| \text{ and } |\sigma(t, x) - \sigma(t, y)| \leq \beta|x - y|$$

for some  $\beta > 0$ . We use the stochastic version of Picard iteration method from ODE. Let  $X_t^{(0)} = x_0$  and define recursively

$$X_t^{(n+1)} = x_0 + \int_0^t m(s, X_s^{(n)}) ds + \int_0^t \sigma(s, X_s^{(n)}) dB_s.$$

For simplicity assume  $0 \leq t \leq 1$ . We want to show that the  $L^2$ -norm is bounded:

$$\mathbf{E} \left[ \left| X_t^{(n+1)} - X_t^{(n)} \right|^2 \right] \leq C \frac{\beta^{2n} t^{n+1}}{(n+1)!}, \quad (**)$$

so that the  $L^2$ -limit  $X_t = \lim_{n \rightarrow \infty} X_t^{(n)}$  exists for each  $t$ . We can choose  $t \in \mathcal{D} \cap [0, 1]$  first and then extend it. Fix  $t$ , we show (\*\*) by induction:

$$\begin{aligned} \text{LHS} &\leq 2\mathbf{E} \left[ \left( \int_0^t \left( m(X_s^{(n)}) - m(X_s^{(n-1)}) \right) ds \right)^2 \right] + 2\mathbf{E} \left[ \left( \int_0^t \left( \sigma(X_s^{(n)}) - \sigma(X_s^{(n-1)}) \right) dB_s \right)^2 \right] \\ &\leq 2\mathbf{E} \left[ \int_0^t \left( m(X_s^{(n)}) - m(X_s^{(n-1)}) \right)^2 ds \right] + 2\mathbf{E} \left[ \int_0^t \left( \sigma(X_s^{(n)}) - \sigma(X_s^{(n-1)}) \right)^2 ds \right] \\ &\leq 4\mathbf{E} \left[ \int_0^t \beta^2 \left| X_s^{(n)} - X_s^{(n-1)} \right|^2 ds \right] \\ &\leq 4\beta^2 \int_0^t C \frac{\beta^{2(n-1)} s^n}{n!} ds = C \frac{\beta^{2n} t^{n+1}}{(n+1)!}. \end{aligned}$$

A particular example of diffusion process is **Bessel process**  $X_t$  which satisfies

$$dX_t = \frac{a}{X_t} dt + dB_t$$

for  $0 \leq t < T = \min\{t : X_t = 0\}$ , where  $a \in \mathbb{R}$  and  $X_0 = x_0 > 0$ .

**Proposition 3.7** (product rule). *Let  $dX_t = R_t dt + A_t dB_t$  and  $dY_t = S_t dt + C_t dB_t$ , then  $d(X_t Y_t) = X_t dY_t + Y_t dX_t + d\langle X, Y \rangle_t$ , where the covariation term*

$$\langle X, Y \rangle_t = \lim_{n \rightarrow \infty} \sum_{j \leq nt} (X_{j/n} - X_{(j-1)/n})(Y_{j/n} - Y_{(j-1)/n}) = \int_0^t A_s C_s ds.$$

*Proof.* Sketch.

$$\begin{aligned} X_{t+\Delta t} Y_{t+\Delta t} - X_t Y_t &= (X_{t+\Delta t} - X_t) Y_{t+\Delta t} + X_t Y_{t+\Delta t} - X_t Y_t \\ &= (X_{t+\Delta t} - X_t) Y_t + (X_{t+\Delta t} - X_t)(Y_{t+\Delta t} - Y_t) + X_t (Y_{t+\Delta t} - Y_t) \\ &\approx Y_t dX_t + d(X_t) d(Y_t) + X_t dY_t \end{aligned}$$

where  $d(X_t)d(Y_t)$  is nonzero. As for the covariation, notice that  $(dB_t)^2 \approx dt$ . □

### In higher dimensions

Let  $B_t = (B_t^1, \dots, B_t^d)$  be a  $d$ -dimensional BM and  $X_t = (X_t^1, \dots, X_t^n)$  a vector process satisfying

$$dX_t^j = R_t^j dt + \sum_{k=1}^d A_t^{j,k} dB_t^k.$$

Since we have  $\langle B^j, B^k \rangle_t = 0$  for  $j \neq k$  and  $\langle B^j, B^k \rangle_t = t$  when  $j = k$ , we have

$$\langle X^j \rangle_t = \sum_{k=1}^d \int_0^t (A_s^{j,k})^2 ds, \quad \text{and} \quad \langle X^j, X^k \rangle_t = \sum_{i=1}^t A_s^{j,i} A_s^{k,i} ds.$$

**Proposition 3.8** (Itô' formula, final form). *Suppose the vector process  $X_t = (X_t^1, \dots, X_t^n)$  satisfies*

$$dX_t^j = R_t^j dt + \sum_{k=1}^d A_t^{j,k} dB_t^k.$$

*Let  $f(t, x) : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}$  be  $C^1$  in  $t$  and  $C^2$  in  $x$ . Then*

$$f(t, X_t) = f(0, X_0) + \int_0^t \partial_t f(s, X_s) ds + \sum_{j=1}^n \int_0^t \partial_j f(s, X_s) dX_s^j + \frac{1}{2} \sum_{j=1}^k \sum_{k=1}^n \partial_{jk}^2 f(s, X_s) d\langle X^j, X^k \rangle_s.$$

*In particular if  $n = d$  and  $X_t = B_t$ , then*

$$df(t, B_t) = \nabla_x f(t, B_t) \cdot dB_t + \left[ \partial_t f(t, B_t) + \frac{1}{2} \Delta_x f(t, B_t) \right] dt$$

## Some applications

Martingale betting strategy, arbitrage: we want the process to reach value 1 at time  $t = 1$ . If it reaches 1 at  $t = 1/2$  we simply stop betting. If it does not reach 1 at  $t = 1/2$ , then we bet more dollars so that the variance is big, and there is a positive probability of reaching 1 at  $t = 3/4$ . Repeat the observation at  $t = 3/4$ . The moral is: in a continuous-time setting we can always "win" within a given time, because we can bet infinitely many times.

European option: Let  $T$  (deterministic) be the call time and  $K$  be the strike price. Suppose the stock (asset) price  $X_t$  is a diffusion process

$$dX_t = m(t, X_t) dt + \sigma(t, X_t) dB_t.$$

The value of the option at time  $T$  is  $F(X_T) = (X_T - K) \vee 0$ . Let  $r(t, x)$  be the interest (bond) rate and

$$R_t = \exp \left\{ \int_0^t r(s, X_s) ds \right\}.$$

One dollar now ( $t = 0$ ) will be promised to value  $R_t$  dollars in the future at time  $t$ . Then

$$\varphi(t, x) = \mathbf{E} \left[ \exp \left\{ - \int_t^T r(s, X_s) ds \right\} F(X_T) \mid X_t = x \right] = \mathbf{E} [(R_t/R_T)F(X_T) \mid X_t = x]$$

is a martingale since

$$M_t = \mathbf{E} [R_T^{-1}F(X_T) \mid \mathcal{F}_t] = R_T^{-1}\varphi(t, x)$$

is a martingale.

## Change of measure

Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be the probability space. Let  $B_t$  be a standard Brownian motion. Suppose that  $M_t$  satisfies  $dM_t = A_t M_t dB_t$  and  $M_0 = 1$ . Then  $M_t = e^{Y_t}$  is a only local martingale, where

$$Y_t = \int_0^t A_s dB_s - \frac{1}{2} \int_0^t A_s^2 ds$$

by Itô's formula. Let  $T_n = \inf\{t : M_t = n \text{ or } \langle M \rangle_t = n\}$  and  $T = \lim_{n \rightarrow \infty} T_n$ . Then  $M_{t \wedge T_n}$  is a nonnegative martingale. Now let  $M_t$  be a nonnegative martingale. Define probability measure  $\mathbb{P}^*$  by  $\mathbb{P}_t^*(V) = \mathbf{E}[M_t \mathbf{1}_V]$  for  $V \in \mathcal{F}_t$ . That is,  $M_t$  is the Radon-Nikodym derivative  $d\mathbb{P}^*/d\mathbb{P}$ .

**Lemma 3.9.** *If one of the following holds, then  $M_s$  is a martingale for  $0 \leq s \leq t$ :*

- (i)  $\mathbf{E}(M_t) = 1$ ;
- (ii)  $\mathbb{P}^*\{T > t\} = 1$ ;
- (iii) (Novikov's condition)  $\mathbf{E}[\exp\{\langle Y \rangle_t/2\}] < \infty$ .

**Theorem 3.10** (Girsanov's theorem). *If  $M_t$  is a nonnegative martingale, then  $dB_t = A_t dt + dW_t$  for  $t < T$ , where  $W_t$  is a standard Brownian motion with respect to  $\mathbb{P}^*$ .*

*Proof.* Firstly,  $W_0 = 0$  and  $t \mapsto W_t$  is continuous. Since  $\mathbb{P}^* \ll \mathbb{P}$  on  $\mathcal{F}_t$ , we have  $\langle W \rangle_t = t$ . Hence by, it suffices to show that  $W_t$  is a  $\mathbb{P}^*$ -martingale.  $\square$

**Remark.** Heuristics of Girsanov's theorem

**Example** (Brownian motion tilted by  $B_t^1$ ). Let  $M_t = B_t$  and  $B_0 = 1$ . Then  $M_t$  is a martingale and  $dM_t = B_t^{-1}M_t dB_t$ . Apply Girsanov's theorem we see that for  $t < T$ ,  $dB_t = B_t^{-1}dt + dW_t$ . Hence  $B_t$  is a Bessel process with  $a = 1$ . Hence for fixed  $t$ , with  $\mathbb{P}^*$ -probability 1,  $0 < B_s < \infty$  for  $0 \leq s \leq t$  and  $T = \infty$ . But with respect to  $\mathbb{P}$ ,  $B_s$  may reach 0. We have tilted the Brownian motion so that the weight of paths reaching 0 is 0.

**Example** (Brownian motion tilted by  $B_t^r$ ). Let  $X_t = B_t^r$  and  $B_t = 1$ . Then  $M_t$  is a martingale and satisfies

$$dX_t = rB_t^{r-1} dB_t + \frac{r(r-1)}{2} B_t^{r-2} dt = \left( \frac{r}{B_t} dB_t + \frac{r(r-1)}{2B_t^2} dt \right) B_t^r$$

by Itô's formula. We use the following trick to get rid of the drift term: let  $M_t = C_t X_t$  where

$$C_t = \exp \left\{ - \int_0^t \frac{r(r-1)}{2B_s^2} ds \right\}.$$

Then use product rule and the fact that  $\langle C, X \rangle_t = 0$  to conclude that  $dM_t = rB_t^{-1}M_t dB_t$ . Apply Girsanov's theorem, we see that  $B_t$  is a Bessel process with  $a = r$ . When  $r \geq 1/2$ ,  $M_t$  is a martingale since  $\mathbb{P}^*\{T = \infty\} = 1$ .

Let  $dX_t = X_t(m dt + \sigma dB_t)$  where  $m = m(t, X_t)$  and  $\sigma = \sigma(t, X_t)$ . We want to find a measure  $\mathbb{P}^*$  under which the drift term disappears and  $X_t$  is a martingale. It suffices to show that  $dB_t = -(m/\sigma) dt + dW_t$ , where  $W_t$  is a standard Brownian motion under  $\mathbb{P}^*$ . Let  $dM_t = (-m/\sigma)M_t dB_t$ . To use Girsanov's theorem, we need that  $M_t = e^{Y_t}$  is a nonnegative martingale.

**Claim:**  $M_t$  is a martingale if  $\mathbf{E}(M_t) = 1$ .

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## Application to finance

### Classification of Lévy process

A process  $\{X_t\}_{t \geq 0}$  is a **Lévy process** if

- (i) it has independent, identically distributed increments, and
- (ii)  $X_t \rightarrow X_0$  is probability as  $t \downarrow 0$ .

For example, Brownian motion  $(m, \sigma^2)$  and Poisson process  $(\lambda)$  are Lévy processes.

An example of Lévy process with discontinuous paths is  $T_a = \inf\{t : B_t = a\}$  for  $B_t$  a standard Brownian motion. It has density  $at^{-3/2}(2\pi)^{-1/2} \exp\{-a^2/2t\}$ . It does not have continuous path, for let  $a = \max\{B_t : t \in [0, 1]\}$ , then  $\lim_{\epsilon \downarrow 0} T_{a+\epsilon} \neq T_a$ . Similarly, let  $B_t = (B_t^1, B_t^2)$  be a standard 2-dimensional Brownian motion. Then  $Y_s = B_{T_s}^1$  is a Lévy process with Cauchy distribution  $s^{-1}\pi(x^2 + s^2)$ , where  $T_s = \min\{t : B_t^2 = s\}$ . It does not have continuous paths. It turns out that Brownian motion is the only Lévy process with continuous paths.

Our goal is to classify all Lévy process.

A random variable  $X$  has **infinitely divisible distribution** if for any  $n \in \mathbb{Z}^+$  we can find a sequence of independent, identically distributed  $Y_1, \dots, Y_n$  such that  $X = Y_1 + \dots + Y_n$  in law.

**Lemma 3.11.** *If  $X_t$  is a Lévy process, then  $X_t$  has an infinitely divisible distribution for each  $t$ . Conversely, for any infinitely divisible distribution  $F$ , there is a Lévy process  $X_t$  such that  $X_1$  has distribution  $F$ .*

*Proof.* Write  $X_t = \sum_{j=1}^n (X_{jt/n} - X_{(j-1)t/n})$  as a telescoping sum. For the converse, □

### Exercise from Wednesday Class (Due this week)

Let  $D \subset \mathbb{R}^2$  be a bounded, simply connected domain containing the origin 0 and  $\hat{D} = D \setminus \{0\}$ . Let  $B_t$  be a standard Brownian motion starting at  $B_0 = x \in \hat{D}$ . Let  $T_s = \inf\{t : |B_t| \leq e^{-s}\}$ ,  $T = \inf\{t : B_t = 0\}$ , and  $\tau = \inf\{t : B_t \notin D\}$ .

(i) Let  $G(x) = \lim_{s \rightarrow \infty} s\mathbb{P}^x\{T_s < \tau\}$ . Show that  $0 < G(x) < \infty$  for  $x \in \hat{D}$ .

Suppose  $D$  is contained in a circle  $C$  of radius  $R$  around the origin, then

$$\mathbb{P}^x\{T_s < \tau\} \leq \mathbb{P}^x\{T_s < \tau_R\} = \frac{\log R - \log |x|}{\log R - \log e^{-s}},$$

where  $\tau_R = \inf\{t : B_t \notin C\}$ . Hence  $\limsup_{s \rightarrow \infty} s\mathbb{P}^x\{T_s < \tau\} \leq \log R + \log(1/|x|) < \infty$ .

On the other hand, suppose  $D$  contains a circle  $c$  of radius  $r$  around the origin, then

$$\mathbb{P}^x\{T_s < \tau\} = \mathbb{P}^x\{\tau' < \tau\}\mathbb{P}^{\frac{x}{2}}\{T_s < \tau\} \geq \mathbb{P}^x\{\tau' < \tau\}\mathbb{P}^{\frac{x}{2}}\{T_s < \tau_r\} = K \frac{\log r - \log(r/2)}{\log r - \log e^{-s}},$$

where  $K = \mathbb{P}^x\{\tau' < \tau\}$ ,  $\tau' = \inf\{t : |B_t| < r/2\}$ , and  $\tau_r = \inf\{t : B_t \notin c\}$ . Hence

$$\liminf_{s \rightarrow \infty} s\mathbb{P}^x\{T_s < \tau\} \geq \liminf_{s \rightarrow \infty} sK \frac{\log r - \log(r/2)}{\log r - \log e^{-s}} = K(\log r - \log(r/2)) > 0.$$

(Monotonicity)

(ii) Show that  $G$  is harmonic in  $\hat{D}$ .

Choose any  $x \in \hat{D}$ . Let  $\epsilon < e^{-s}$  and  $\epsilon < |x|$ . Then

$$s\mathbb{P}^x\{T_s < \tau\} = s \int_{|z-x|=\epsilon} \mathbb{P}^z\{T_s < \tau\} ds$$

satisfies the mean-value property, where  $s$  is the normalized surface measure. By monotonicity from part (i), we can exchange limit and integral. Hence  $G(x) = \text{MV}(G, x, \epsilon)$  satisfies the mean-value property, and is harmonic in  $\hat{D}$ .

(iii) Show that if  $x_n \rightarrow z \in \partial D$ , then  $G(x_n) \rightarrow 0$ .

(A conformal invariance argument)

(iv) Let  $M_t = G(B_t)$  for  $0 \leq t < t \wedge T$ . Show that  $M_t$  is a local martingale for  $t < T \wedge \tau$ .

Use Itô's formula:  $dM_t = dG(B_t) = \nabla G(B_t) \cdot dB_t + \Delta G(B_t) dt/2 = \Delta G(B_t) \cdot dB_t$ , since  $G$  is harmonic by part (ii). Hence  $M_t$  is a local martingale.

(v) Use Girsanov's theorem to find a new probability measure under which  $T < \infty$  and  $T < \tau$ .

(vi) Is  $M_t$  a martingale?



Let  $\{M_t\}_{t \geq 0}$  be a nonnegative discrete-time martingale with  $M_0 = 1$ . Let  $M_\infty = \lim_{n \rightarrow \infty} M_n$ . Let  $\mathcal{F}_n$  be the  $\sigma$ -algebra generated by  $\{M_0, \dots, M_n\}$ . Define measure  $\mathbb{Q}_n$  on  $\mathcal{F}_n$  by  $\mathbb{Q}_n(V) = \mathbf{E}[M_n \mathbf{1}_V]$  for  $V \in \mathcal{F}_n$ . Let  $\mathbb{Q}$  be the measure that is equal to  $\mathbb{Q}_n$  when restricted to  $\mathcal{F}_n$ . Show that if  $\mathbb{Q}\{M_\infty < \infty\} = 1$  or  $\mathbb{Q}\{\sup_n M_n < \infty\} = 1$ , then  $\mathbf{E}[M_\infty] = 1$ .